

# Strategic Thinking Skills: A Key to Collective Economic Success

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*We conduct a large-scale experiment to measure elementary aspects of strategic thinking skills and their linkage to labor market outcomes. Two incentivized measures of higher-order rationality and backward induction are developed. Males' (females') strategic thinking skills are positively (negatively) associated with individual labor income. However, among married individuals, strategic thinking skills are significantly and positively associated with their household labor income regardless of gender, highlighting the importance of strategic thinking skills for collective economic success. We argue that the intrahousehold channels encompassing collective labor supply with home-to-workplace spillover and marriage assortative matching offer the most plausible explanation for our findings.*

*JEL: C91, D91, J24*

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Humans, as social creatures, engage in numerous interpersonal interactions throughout their lives. The ability to understand motivations, anticipate the behavior of other people, and respond to others is essential to social relationships and economic success. People with higher strategic thinking skills may maintain better interpersonal relations and secure higher economic returns.

In this paper, we argue that strategic thinking is a multi-dimensional skill of significant economic importance, distinct from the traditional collection of cognitive and noncognitive skills. To measure skills of strategic thinking and support our

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claim, we resort to the experimental methods of detecting the depth of strategic reasoning in situations of simultaneous and sequential interaction. We conduct a large-scale experiment and investigate whether these measures of strategic thinking skills are associated with labor market outcomes.

The canonical approach to understanding interpersonal interaction is based on the concept of Nash equilibrium, which requires no limit in human abilities of strategic reasoning (Aumann and Brandenburger, 1995; Polak, 1999). While the standard equilibrium approach offers a powerful tool for analyzing strategic interactions, it overlooks the possibility that individuals differ in their capability of strategic reasoning, particularly when there is no opportunity to learn from repeated play. In contrast, the experimental economics literature has documented that human reasoning in interpersonal interactions in laboratories and fields is far below the level of sophistication assumed by the standard theory and exhibits a large degree of individual heterogeneity (e.g., Nagel, 1995; Camerer, Ho and Chong, 2004; Crawford, Costa-Gomes and Iriberri, 2013).

We consider two elementary aspects of strategic reasoning in interpersonal interaction: (i) engaging in introspective reasoning in situations of simultaneous interaction, and (ii) exercising anticipatory reasoning or backward induction in situations of sequential interaction. These elements are essential and mutually distinct in the game-theoretic analysis of strategic interaction: the former relates to higher-order rationality or rationalizability (Bernheim, 1984; Pearce, 1984), while the latter is key to sequential rationality serving as a refinement of the Nash equilibrium (Selten, 1965). Although our two measures are unlikely to capture the entire spectrum of strategic thinking skills, our attempt is still meaningful as the first step to investigate the relationship between individuals' strategic thinking skills and economic outcomes.

To measure higher-order rationality (HOR), we develop a five-person line-network game, motivated by Kneeland (2015). A series of two-person normal-form games are connected by a network structure of opponents. Participants make a request for money (Arad and Rubinstein, 2012) in each of the five different positions in a random order without any feedback. The choice data from the line-network structure of the opponent enable us to measure different levels of HOR reasoning. Our second measure is developed based on a series of two-player sequential-move games that require different steps of backward-induction (BI) reasoning (Dufwenberg, Sundaram and Butler, 2010; Gneezy, Rustichini and Vostroknutov, 2010). Each game has a first-mover advantage. Human participants always move first and play against a computer player, programmed to play an optimal strategy. This design allows us to measure BI reasoning at the individual level.

To implement these measures, we recruited participants from the Singapore Life Panel (SLP), a nationally representative sample of people 50–70 years old in Singapore. The final sample consists of 2,146 Singaporeans whose age ranges between 50 and 65. We take advantage of detailed information on their socioeconomic characteristics as well as a rich set of cognitive and noncognitive skill

measures available in the large-scale panel data. To demonstrate the robustness of the descriptive statistics of our strategic thinking measures to a younger population, we also recruited 786 participants from the Korean Labor and Income Panel Study (KLIPS), a nationally representative sample of urban households and individuals in South Korea.

We view our measures of strategic reasoning as capturing fundamental aspects that all types of real-world strategic interaction commonly share. Those with higher capabilities in exercising strategic reasoning are more likely to better serve their social and economic interests over their life span. The HOR measure involves introspective reasoning of others' behavior by placing oneself in others' situations, whereas the BI measure concerns a comprehensive evaluation of multiple future contingencies in order to arrive at a present decision. Hence, the relevance of one measure over the other in understanding human behavior may depend on the nature of strategic interaction. To examine these conjectures, we estimate the relationship between strategic thinking skills and labor market outcomes at the individual and household levels. In particular, we examine the associations separately by gender following the literature documenting gender differences in labor supply decisions within the couple (e.g. [Altonji and Blank, 1999](#); [Kuziemko et al., 2018](#)).

We find strong gender-dependent associations between individuals' labor income and their strategic thinking skills. Male labor income is *positively* associated with strategic thinking skills. These associations are robust to conditioning on a conventional set of cognitive and noncognitive skills as well as sociodemographic characteristics: a one-level increase in the BI score and a one-standard-deviation (SD) increase in the HOR score are significantly associated with 37 percent and 58 percent *increases* in own labor income, respectively. Regarding female labor income, we find the opposite pattern; female respondents' BI score is *negatively* associated with their individual labor income. This result is robust to conditioning on cognitive and noncognitive skills as well as sociodemographic characteristics: a one-level increase in female respondents' BI score is significantly associated with a 47 percent *decrease* in their labor income.

To gain a deeper understanding of the gender-dependent association between individual labor income and strategic thinking skills, we investigate the *extensive margin of individual labor supply* (i.e., whether one earns positive labor income). We find strong gender-dependent associations between labor supply and strategic thinking skills. Male respondents with higher HOR and BI scores are significantly more likely to work and less likely to be retired or unemployed. In contrast, female respondents with higher BI scores are less likely to work and, in the composition of labor status, more likely to be a homemaker and thus out of the labor force. The observed gender-dependent patterns suggest that strategic thinking skills are associated not only with individual decision making of labor supply but also with intrahousehold interactions. In fact, there is a strong positive linkage between female participants' skills of strategic thinking and their spouses' labor income,

suggesting that a variety of intrahousehold interactions, including partner matching, intrahousehold labor supply decision making, and spillover/crossover between home and work, can play a role.

We then explore how *household labor income for married respondents* is associated with their individual strategic thinking skills. We find *gender-independent positive* associations: a one-SD increase in the HOR score is associated with an approximately 60 percent increase for males and females. Restricting attention to the non-working female married sample, we find that a one-level increase in the BI score and a one-SD increase in the HOR score are significantly associated with a 87 percent increase and a 114 percent increase in female household labor income, respectively. These gender-independent positive associations with household labor income for married respondents suggest the potential role of strategic thinking skills in collective economic success at the household level.

How can we reconcile the observed household-level, gender-independent, and individual-level, gender-dependent associations between strategic thinking skills and labor market outcomes? To answer this question, we consider several plausible channels. First, we consider the *labor market channels* through which individuals with higher strategic thinking skills contribute better to team production in the workplace (*workplace channel*) or are sorted into professions and industries with higher demand for these skills (*occupation-choice channel*). While the workplace channel can explain the positive association for males' individual labor income, it is difficult to rationalize the opposite pattern found for females' individual labor income, gender-dependent associations for the extensive margin labor supply, and gender-independent associations for household labor income. Regarding the occupation-choice channel, we conduct a regression analysis on whether respondents' occupation choices are associated with strategic thinking skills and find little evidence supporting this explanation.

Given the challenges in rationalizing the empirical findings solely through labor market channels, we direct our focus towards the *intrahousehold interaction channels*. Two plausible channels emerge: 1) collective labor supply with home-to-workplace spillover and 2) assortative matching on strategic thinking skills in marriage. The first channel suggests that when the allocation of household and workplace production tasks between married individuals occurs based on their comparative advantage or gender norms, strategic thinking skills can facilitate coordination and enhance the spillover effect from home to the workplace (e.g., [Apps and Rees, 1997](#); [Chiappori, 1997](#); [Benham, 1974](#); [Barnett, 1994](#)). The second channel proposes that individuals tend to marry partners with similar strategic thinking skills.

We contend that these two intrahousehold interaction channels complement each other in explaining our data. While the former channel accounts for the essential empirical patterns of gender-dependent associations with individual labor income and gender-independent associations with household labor income, it does not address the positive correlation of strategic thinking skills between married

individuals found in our data. Conversely, the latter channel can explain the positive correlation of strategic thinking skills between spouses and certain patterns of the associations between strategic thinking skills and labor income. However, this channel alone cannot explain why certain couples specialize in household production and workplace production, nor why individual income resulting from this specialization exhibits gender-dependent correlations with strategic thinking skills. Only by combining these two channels can we adequately account for the key empirical findings observed in this study.

One potential concern in the empirical analysis pertains to the measurement errors in eliciting strategic thinking skills and the absence of corrections for any noise in these elicitation methods when establishing the relationship between strategic thinking skills and labor outcomes. To address this concern, we adopt the method of obviously related instrumental variables (ORIV) proposed by [Gillen, Snowberg and Yariv \(2019\)](#). It is important to note that, in implementing the ORIV approach, we make a compromise by assuming that the BI and HOR measures capture the same underlying trait with errors, departing from our key assumption that strategic thinking is a multi-dimensional construct and our HOR and BI measures capture two distinct facets of strategic thinking skills. Nevertheless, our analysis reveals that the ORIV method generates estimation results consistent with the gender-dependent associations with individual labor income and the gender-independent associations with household labor income. Hence, we argue that our primary findings are unlikely to be compromised by concerns regarding measurement errors.

The existence of a gender gap in labor market outcomes is well-documented in the literature ([Goldin, 1990](#); [International Labour Organisation , ILO](#)), and the primary objective of this paper is not to examine or identify the exact sources of the observed gender gap. Instead, we acknowledge the gender gap as a given fact and focus on establishing gender-dependent associations between strategic thinking skills and labor market outcomes. We argue that these associations are driven by intrahousehold interaction channels, including assortative matching based on strategic thinking skills in marriage, as well as the facilitation of intrahousehold task specialization by strategic thinking skills.

Recently, researchers have explored the relationship between strategic thinking skills and other cognitive and noncognitive skills outside a controlled laboratory environment. In repeated strategic interactions, [Gill and Prowse \(2016\)](#) find that both cognitive ability and noncognitive skills are correlated with level- $k$  thinking. Using a sample of children aged 5–12 years old, [Fe, Gill and Prowse \(2022\)](#) find that higher theory-of-mind and cognitive abilities predict a higher degree of strategic thinking skills in competitive games. To our knowledge, we are the first paper that provides empirical evidence that strategic thinking skills are important components shaping one's individual and collective economic success, even after controlling for a variety of individual characteristics, including educational attainment, family background, cognitive skills, and noncognitive skills.

We also contribute to the literature on human capital and its importance in the labor market. While human skills are multidimensional in nature (e.g., Heckman, Stixrud and Urzua, 2006b; Cunha and Heckman, 2007), the literature has traditionally focused on cognitive skills (e.g., Herrnstein and Murray, 1994; Hanushek and Woessmann, 2008) and, more recently, on a growing list of noncognitive skills, including personality, grit, and self-efficacy (Bowles, Gintis and Osborne, 2001; Heckman and Rubinstein, 2001; Borghans, Ter Weel and Weinberg, 2008; Almlund et al., 2011; Lindqvist and Vestman, 2011; Heckman, Jagelka and Kautz, 2019). A few studies document the importance of social skills in the labor market, related to the theme of this paper. Deming (2017) reports that the U.S. labor market increasingly rewards social skills by providing higher wages for jobs requiring high levels of social interaction. Borghans, Ter Weel and Weinberg (2008) document that sociability in youth is a good predictor of later job assignment and that the returns to interpersonal styles vary across jobs depending on the types of interpersonal tasks. Conti et al. (2013) use friendship nomination in high school as a proxy of social skills and show that it is a good predictor of future earnings. Our study uses game-theoretic measures of an individual’s strategic thinking skill, uncovers the gender-dependent associations between these measures and individual labor incomes.

The remainder of the paper is structured as follows. Section I describes how we measure strategic thinking skills and presents the empirical features of these measures. Section II discusses how our measures of strategic thinking skills are correlated with each other, and with other standard cognitive and noncognitive measures. In Section III, we report the estimation results documenting the economic importance of strategic thinking skills. Section IV considers a few plausible channels to rationalize our key empirical findings. We check the robustness of the baseline findings in Section V. Section VI concludes the paper.

## I. Experimental Design and Measurements

### A. Measuring Strategic Thinking Skills

LINE(-NETWORK) GAME: HIGHER-ORDER RATIONALITY MEASURE. — The HOR measure is built on the *Line-Network Game* (hereafter, the Line Game), developed to identify individual heterogeneity in conducting introspective thinking during the iterative elimination of strictly dominated strategies (i.e., rationalizability, Bernheim (1984) and Pearce (1984)). It is a five-person simultaneous-move, dominance-solvable, network game, à la Kneeland (2015). It consists of a series of two-person games, each of which is adapted from the 11-20 money request game of Arad and Rubinstein (2012), with the opponent structure determined by the line network. Figure A1 in Appendix A presents a sample decision screen for the Line Game.

In the game, there are five positions—A, B, C, D, and E. Each player is assigned

to one of the five positions and makes a decision *simultaneously* and independently; the position A player makes a money request of either S\$10 (US\$7.38) or S\$50 (US\$36.9); players in any other position make a money request from 5 options: S\$10, S\$20, S\$30, S\$40 or S\$50.<sup>1</sup> The payoff of the position A player is the amount of money s/he requested. The payoffs of players in any other position consist of two parts: each player receives 1) the amount of money s/he requested and 2) an additional amount of S\$100 if and only if the money s/he requested is S\$10 lower than the money requested by his/her opponent. The opponent of each player is defined as a player who occupies a position to the left of that player in the line network.<sup>2</sup>

Each individual plays the game five times, in a random order, in each of the five positions. The following set of choices is implied by the full rationality of players: First, the player in position A chooses S\$50. Correctly anticipating this choice, the player in position B chooses S\$40. This iterative process continues, resulting in the choices of S\$30, S\$20, and S\$10 by the players in positions C, D, and E, respectively.

Our goal is to measure how well each individual in this simultaneous-move environment performs introspective thinking by forming a correct belief about the choices made by others who are not necessarily fully rational.<sup>3</sup> To obtain this measurement, we define “HOR score” as the average expected payoff of an individual, calculated based on his/her choice in each position matched with the empirical distribution of his/her opponent’s choices. Specifically, we first obtain the empirical choice distribution for each position from our choice data. We then match an individual’s choice in each position with the empirical choice distribution for his/her opponent’s position to obtain the expected payoff for each position. Finally, we take the average of the expected payoffs from all five positions to obtain the HOR score of the individual. A standardized version of this measure (having a mean of 0 and a standard deviation of 1 by gender) will be used for our main empirical analyses. In addition, two other variables—discretized HOR score and HOR order measure—are constructed based on individuals’ performance in the Line Game and used to demonstrate the robustness of our results. These two variables will be defined and discussed in Appendix [A.A16](#).

Table 1 below presents the empirical distributions of the HOR scores obtained from the SLP and KLIPS data. The two distributions share the same qualitative features, although the distribution from the SLP data first-order stochastically dominates that from the KLIPS data slightly.<sup>4</sup>

<sup>1</sup>S\$1 is equivalent to 0.74 US\$ or 0.66 euro as of January 29, 2022.

<sup>2</sup>The position A player is the opponent of the position B player. The position B player is the opponent of the position C player. The position C player is the opponent of the position D player. The position D player is the opponent of the position E player. However, the opponent relationship is asymmetric. For example, the position B player is *not* the opponent of the position A player.

<sup>3</sup>The full-rationality benchmark is also considered, and our results do not depend on which measure we adopt. For more details, see Appendix [A.A16](#).

<sup>4</sup>The distribution of the HOR scores by gender is reported in Appendix [A.A2](#).



Table 1—: Distribution of the HOR scores

| Data  | Min  | 10th | 25th | 50th | 75th | 90th | Max  |
|-------|------|------|------|------|------|------|------|
| SLP   | 23.9 | 38.0 | 41.5 | 49.4 | 65.8 | 68.6 | 72.6 |
| KLIPS | 21.5 | 36.1 | 40.3 | 47.8 | 54.5 | 63.9 | 67.6 |

LIFT GAME: BACKWARD INDUCTION MEASURE. — The BI measure is built upon the *Lift Game*, a two-person sequential-move game that has been considered in the literature, e.g., the Race Game (Gneezy, Rustichini and Vostroknutov, 2010) and the Game of 21 (Dufwenberg, Sundaram and Butler, 2010). The game is developed to identify individual heterogeneity in conducting BI reasoning (Selten, 1965).

To describe the game, imagine that two players—player 1 (she) and player 2 (he)—get on the same lift on the ground floor (floor 0). Player 1 first decides how many floors to go up. She has three choices: one, two, or three floors up. Then, the lift moves up to the floor chosen by the first mover. Player 2 next decides how many floors to go up while facing the same three choices. There is a predetermined and publicly known target floor, denoted by  $k > 1$ . The two players take turns moving until the lift arrives at the target floor. The player who presses the target floor button  $k$  in the lift wins the game.

There are four rounds of the Lift Game with different target numbers. Human participants, always being the first mover, play the four rounds of the Lift Game in a random order against a computer player. The computer player is programmed to choose an optimal move in every information set while uniformly randomizing whenever indifferent. This experimental design ensures that the human player does not face any strategic uncertainty about what the opponent will do. Thus, the design faithfully implements an environment where rationality is common knowledge. The target numbers we chose are 5, 11, 14, and 17. Game  $G_k$  denotes the Lift Game with target number  $k$  in which  $k \in \{5, 11, 14, 17\}$ . Figure A2 in Appendix A presents sample decision screens for  $G_{14}$ .

Each game  $G_k$  has a winning strategy for the first mover. For example, the winning strategy for the first mover in  $G_5$  is to go up to the 1st floor in her first turn and then to the 5th floor in her second turn. The winning strategy for each game is presented in the last column of Table 2. The existence of a winning strategy for the first mover implies that an individual who can perform two steps of backward inductive reasoning should win game  $G_5$ . Similarly, an individual who can perform three, four, and five steps of backward inductive reasoning should win games  $G_{11}$ ,  $G_{14}$ , and  $G_{17}$ , respectively. Hence, we assign the BI score of 1 to an individual who did not win game  $G_5$ , the BI score of 2 to an individual who won  $G_5$  but not  $G_{11}$ , and so on. Table 2 below illustrates



the classification criterion we used for the BI score.<sup>5</sup> The last two rows present the empirical distribution of the BI scores. For example, in the SLP sample (our baseline Singapore data), 41.7% and 34.7% of participants won the game only once and twice out of the five rounds, respectively. Those who won each of the five rounds account for only 1.9% of the participants. We also find the similar pattern in the Korean sample (KLIPS data) showing that 40.1% and 36.8% of the participants won only once and twice out of the five rounds, respectively. Despite the population difference, it is notable that the distribution of the BI score is fairly consistent between the two countries. As a robustness check, we consider alternative measures—categorical BI score and BI counting score—in Appendix A.A16.

Table 2—: Classification and distribution of the BI score

| Score    | 1     | 2     | 3     | 4    | 5    | Winning Strategy |
|----------|-------|-------|-------|------|------|------------------|
| $G_5$    | Lose  | Win   | Win   | Win  | Win  | 1-5              |
| $G_{11}$ | -     | Lose  | Win   | Win  | Win  | 3-7-11           |
| $G_{14}$ | -     | -     | Lose  | Win  | Win  | 2-6-10-14        |
| $G_{17}$ | -     | -     | -     | Lose | Win  | 1-5-9-13-17      |
| SLP      | 41.7% | 34.7% | 19.7% | 2.1% | 1.9% | -                |
| KLIPS    | 40.1% | 36.8% | 17.6% | 1.7% | 3.9% | -                |

As noted by Alaoui and Penta (2016) and Jin (2021), an individual’s depth of reasoning may depend not only on their ability to conduct rounds of introspection (in the case of our HOR measure) or backward induction (in the case of our BI measure) but also on their beliefs about their opponents. It is important to highlight that our experimental design for the BI measure involves human participants playing against a computer player programmed to use an optimal strategy on every path, leaving no room for beliefs about the opponent’s strategic thinking skills to influence choices. For our HOR measure, we address the concern in Appendix A.A16 by demonstrating the robustness of our results to various definitions of the HOR measure.

### B. Data and Procedures

DATA. — For the main empirical analysis, we recruited our study participants from the Singapore Life Panel (SLP), a nationally representative internet-based monthly panel survey in Singapore. Most respondents were 50–70 years old when

<sup>5</sup>Our identification method only captures the *upper bound* of the BI reasoning steps an individual could perform. For example, a person who can perform only two steps of BI reasoning could randomly make a first move that coincidentally matches the winning strategy in  $G_{11}$ .

it was launched in July 2015. The SLP has been collecting a rich array of individual and household characteristics, such as family structure, labor market outcomes, and health. The online nature of the survey allows researchers to flexibly ask various types of questions in an interactive manner.

In the August 2017 wave, we invited 3,595 respondents aged 50–65 years to participate in our study. We deliberately decided to not invite those aged over 65 years to focus on the working-age population. At the time of the survey, the Retirement and Re-employment Act in Singapore mandated most employers offer continued employment until 65. In addition, the official pension claiming age in Singapore called the Payout Eligibility Age, is 65. Participants were informed that they would receive S\$5 upon completing the tasks in our study and up to S\$150 based on their performance in each task. A total of 2,787 (78%) accepted our offer, and 2,146 completed all tasks in our study.

We also recruited 786 participants from the Korean Labor and Income Study (KLIPS), a nationally representative sample of urban households and individuals in South Korea. Because the KLIPS sample covers younger ages below 50, it can demonstrate the robustness of the descriptive statistics of strategic thinking measures with respect to age despite the small sample size.

As measures of cognitive ability, we use educational attainment and two internationally popular and well-validated tests of fluid intelligence and social cognition. The Intelligence Structure Test (IST) is our measure of fluid intelligence (Cattell, 1963). It is an internationally used and popular *nonverbal* cognitive ability test first developed in 1953 (Beauducel et al., 2010). The validity and reliability of the IST as a measure of cognitive ability have been established over more than 1800 samples. The figural matrix part of the IST consists of 20 questions and is very similar to the Raven’s Matrices test. Figure A3 in Appendix A shows a sample question.<sup>6</sup>

The Reading the Mind in the Eyes Test (hereafter, Eyes Test), developed by Baron-Cohen et al. (2001), is our measure of an individual’s theory of mind or social cognition, i.e., an individual’s ability to recognize another individual’s mental state (Astington, Harris and Olson, 1988). It concerns reading cues in face-to-face human interaction, ignored in mathematical descriptions of strategic interaction but found to play an important role (e.g., Scharlemann et al., 2001; Stirrat and Perrett, 2010). The Eyes Test contains 28 questions, each of which shows a photo of the human eye area, and asks the respondent to choose a word that best describes the person’s mental state. Figure A4 in Appendix A shows a sample question. The validity and reliability of the Eyes Test are also well established across many countries (Olderbak et al., 2015).

As measures of noncognitive traits and economic preferences, we use financial planning horizon, risk tolerance, self-efficacy, and personal optimism. The definitions of the cognitive and noncognitive trait variables are included in Appendix

<sup>6</sup>The full-length test includes other dimensions of intelligence such as verbal memory and numerical knowledge, but we could not implement those components due to the survey time constraints.

#### A.A4.

Column (1) of Table A3 in the Appendix A.A4 reports sample characteristics of the study participants who completed all tasks. Participants are, on average, 58.5 years old, and almost half of them are male; 91 percent are ethnic Chinese, and 82 percent are married with almost 3 children. Of the participants, 45 percent received at least postsecondary education, and the average cognitive ability, in terms of the IST score, is 10.8 (out of 20), which is more than 1 point higher than 9.6, the score corresponding to an IQ score of 100 according to the test's manual (Beauducel et al., 2010). Of the participants, 70 percent reported positive annual labor income, earning approximately S\$50,283 on average. We do not have information on hourly wages due to a lack of data on specific work hours.

Column (2) of Table A3 reports the sample characteristics of study dropouts, i.e., those who accepted our invitation but did not complete the survey module. Column (3) of Table A3 presents the sample characteristics of nonparticipants, i.e., those who did not accept our invitation. In general, participants, dropouts, and nonparticipants are similar in terms of individual characteristics. Table A5 shows the descriptive statistics of the KLIPS sample. For the regression analysis, we focus on the SLP sample and do not use the KLIPS sample due to the small sample size.

PROCEDURES. — Our study comprised two tasks that correspond to the strategic thinking measures discussed in Section I.A. In Task I, each participant played four rounds of the Lift Game. In Task II, each participant was randomly matched with four other participants and played five rounds of the Line Game.<sup>7</sup>

The cash payment consisted of three parts. First, upon completing the experiment, every participant received the show-up fee of S\$5. Second, for each participant, one game (out of 4 games) in Task I was chosen randomly; the winning participant received S\$5, and the others received S\$0. Third, the dollar amount each participant earned in one randomly chosen round (out of 5 rounds) in Task II was paid to the participant depending on the outcome of a lucky draw in which each participant had a 10 percent chance of winning. Given the uncertainty of the lucky draw for actual earnings in Task II, we increased the money stake in Task II so that the expected earnings from Task II are similar to those from Task I. Participants received a minimum amount of S\$5 (US\$3.7) and a maximum amount of S\$150 (US\$110.6) by participating in the experiment, which lasted approximately twenty minutes on average.<sup>8</sup>

<sup>7</sup>In Tasks I and II, after reading the instructions, participants were asked to answer a few comprehension quiz questions and to play a practice round. The scripts for the experimental instructions are available in Appendix A.A1. The Eyes Test was conducted after these two tasks.

<sup>8</sup>Due to an administrative restriction, the payment was delivered in the following month in the form of a cash voucher for the largest grocery store chain in Singapore.

## II. Correlations Among Our Measures

Table 3 provides the Pearson correlation coefficients and statistical significance levels between the strategic thinking measures obtained from the SLP and KLIPS datasets. It also presents the correlation coefficients between our strategic thinking measures and other cognitive/noncognitive skills.

Upon examination, several observations become evident. Firstly, our strategic thinking measures exhibit positive correlations with conventional measures of cognitive and noncognitive skills. Specifically, both the HOR and BI measures show correlations with the Eyes test score and IST score. However, it is important to note that the degree of correlation is not excessively large, comparable to some correlations between different elicitation methods of common economic preferences reported in [Chapman et al. \(2023\)](#). This positive correlation aligns with the conceptual connection of our strategic thinking measures to the theory of mind (in the case of the Eyes test) and intellectual/cognitive ability (in the case of the IST). Nevertheless, we emphasize that our strategic thinking measures are fundamentally distinct from the theory of mind and cognitive skills for two reasons. Firstly, none of the known cognitive skill measures involve interactive decision-making. Secondly, the theory of mind, as measured by the Eyes test, primarily captures an individual’s ability to perceive and interpret facial cues in interpersonal interactions but not any kind of higher-order reasoning.

Table 3—: Correlations among strategic thinking skill measures and other skills

| SLP                        |          |           | KLIPS                   |          |           |
|----------------------------|----------|-----------|-------------------------|----------|-----------|
|                            | BI Score | HOR Score |                         | BI Score | HOR Score |
| HOR Score                  | 0.316*** |           | HOR Score               | 0.127*** |           |
| Eyes Test Score            | 0.116*** | 0.161***  | Eyes Test Score         | 0.183*** | 0.048     |
| Postsecondary education    | 0.013    | 0.064***  | Postsecondary education | 0.032    | 0.013     |
| IST Score                  | 0.230*** | 0.213***  | Risk Tolerance          | -0.011   | -0.045    |
| Financial planning horizon | 0.029    | 0.036*    | Patience                | -0.208   | 0.024     |
| Risk Tolerance             | 0.005    | 0.026     | Openness                | 0.061*   | 0.028     |
| Self-efficacy              | 0.040*   | 0.055***  | Conscientiousness       | -0.0004  | 0.078**   |
| Personal Optimism          | 0.077*** | 0.123***  | Extraversion            | 0.074**  | 0.081**   |
|                            |          |           | Agreeableness           | 0.029    | 0.052     |
|                            |          |           | Neuroticism             | 0.065*   | 0.013     |

Note: \* denotes statistical significance at 0.10; \*\* at 0.05; \*\*\* at 0.01.

Secondly, our two strategic thinking measures exhibit a correlation with each other. The correlation coefficients are 0.316 in the SLP data and 0.127 in the KLIPS data. This positive correlation is expected due to their conceptual relationship. However, we assert that strategic thinking skills are multi-dimensional, and our measures of HOR and BI capture distinct aspects of strategic thinking skills. The HOR measure, derived from the money request game, assesses introspective thinking ability by evaluating an individual’s capacity to think strategically on behalf of others, empathetically placing oneself in their position. Conversely, the

BI measure, derived from the lift game, focuses on anticipatory thinking ability, requiring individuals to consider various potential future scenarios to inform decision-making in the present moment. In summary, the BI measure necessitates a comprehensive evaluation of multiple future contingencies to arrive at a present decision, while the HOR measure involves placing oneself in the positions of up to four different players. Consequently, the HOR measure and the BI measure aim to capture two distinct aspects of strategic thinking skills.<sup>9</sup>

The utilization of separate measures to capture different dimensions of a particular skill or ability is well-accepted in the existing literature. For instance, the Wechsler Adult Intelligence Scale, one of the most widely-used measures of intelligence quotient, consists of several performance subtests that assess distinct dimensions of cognitive ability.

To investigate the possibility of assortative matching on strategic thinking skills in marriage, we examine the correlations of strategic thinking skills within married couples who both participated in our study. We find substantial correlations for strategic thinking skills, with a coefficient of 0.44 for the BI score and 0.45 for the HOR score. However, it is worth noting that we also observe similarly high correlations for traditional cognitive and non-cognitive skills within the couples. Specifically, the correlations are 0.46 for education, 0.73 for the IST score, 0.53 for the Eyes test score, 0.53 for subjective risk preference, and 0.50 for self-efficacy.

### III. Strategic Thinking and Labor Market Outcomes

To establish that strategic thinking skills are strong predictors of an individual's economic outcomes, we consider the individual and household labor incomes of participants as the real-world outcomes of interest. Identifying the determinants of individual- and household-level labor income is a key area of research in labor economics (e.g., [Mincer, 1958](#); [Pencavel, 1986](#); [Chiappori, 1988](#); [Miles, 1997](#); [Blundell and Macurdy, 1999](#); [Heckman, Stixrud and Urzua, 2006a](#)). The literature has shown that various skills contribute to inequality in labor market outcomes ([Heckman, 1995](#); [Katz and Autor, 1999](#); [Heckman and Kautz, 2012](#)). Therefore, it is natural for us to investigate whether our measures of strategic thinking skills can independently explain variations in labor income.

We use annual labor income data collected in January 2015 (i.e., annual labor income earned during the calendar year 2014) for the main empirical analysis.<sup>10</sup> The participants in our study are aged 50–65 years; thus, 30% of participants

<sup>9</sup>We do not claim that these two elementary aspects encompass the entirety of strategic thinking skills. There are additional strategic thinking skills that are not captured by our measures, including (but not limited to) 1) learning ability, which refers to how quickly individuals can acquire knowledge about others' strategies, and 2) belief updating, as described in the concept of cursed equilibrium by [Eyster and Rabin \(2005\)](#). Our intent is to contribute to the understanding of strategic thinking by focusing on elementary aspects captured by the HOR and BI measures, recognizing that there are broader and more nuanced facets of strategic thinking that warrant further investigation.

<sup>10</sup>We extend the analysis by pooling multiyear income data in Section [A.A13](#). The results remain robust.

report zero annual labor income. Hence, it is also important to study the association between strategic thinking skills and the extensive margin of labor supply. We define the extensive margin as a binary indicator that takes the value of 1 if a participant has a positive labor income and 0 otherwise. The extensive margin analysis refers to the analysis of labor income using this binary indicator as a dependent variable, while the intensive margin analysis refers to the analysis omitting zero-income earners.

We conduct the empirical analysis separately by gender. The gender gap in labor income and labor market participation has historically been substantial and persistent, although it has decreased to some extent over recent decades (Altonji and Blank, 1999; Kuziemko et al., 2018). This gap depends on the degree of gender discrimination in hiring and workplace relations as well as on differences in gender roles in intrahousehold labor supply (Chiappori, 1992; Fernandez and Fogli, 2009; Bertrand, Kamenica and Pan, 2015; Charles, Guryan and Pan, 2018). Unless such gender differences in the labor market are orthogonal to strategic thinking skills, establishing their association with labor outcomes would be biased if we pool the data over gender. We report the mean and 95 percent confidence intervals of annual labor income (including zero income), the likelihood of working (i.e., reporting a positive annual labor income), and annual labor income (excluding zero income) by the ranking of each measure of strategic thinking skills and cognitive ability in Figures A6, A7, and A8 in Appendix A, respectively.

We transform the annual labor income variable with the inverse hyperbolic sine (IHS) function (Burbidge, Magee and Robb, 1988) and use it as the primary dependent variable in the regression analysis. The IHS transformation has the same interpretation of the log transformation (i.e., percent change) but provides the advantage that it is defined at zero. Thus, we do not need to drop zero-income earners from the sample. This transformation method has been widely used in the literature analyzing medical spending, wealth, and savings (Carroll, Dynan and Krane, 2003; Pence, 2006; Gelber, 2011) as well as earnings (Powell, 2020), in which the variable of interest frequently takes the value of zero.

#### A. Strategic Thinking and Individual Labor Outcomes

Our analysis proceeds in three steps. First, we run a baseline regression of the IHS transformation of an individual's labor income on a measure of strategic thinking skills while controlling for sociodemographic variables. These controls include age group dummies, gender, ethnicity, marital status, number of children, spouse's age, and a dummy for a missing observation of spouse's age (mostly for unmarried respondents).

Second, we additionally control for educational attainment and cognitive abilities measured by the IST score and the Eyes Test score. Educational attainment and cognitive abilities are traditionally considered major determinants of labor income (Becker, 1964; Mincer, 1975; Heckman, Stixrud and Urzua, 2006b). Thus,

we control for them in assessing the robustness of the association between strategic thinking skills and labor income.<sup>11</sup>

Third, we further control for noncognitive skills and economic preferences available in the SLP data—risk tolerance, financial planning, self-efficacy, and personal optimism—following the literature documenting the role of noncognitive skills and preferences in economic outcomes (Heckman, Stixrud and Urzua, 2006*b*; Almlund et al., 2011; Falk et al., 2018; Heckman, Jagelka and Kautz, 2019). In addition, we include a response time taken to complete a corresponding experiment and the random order of tasks in the experiment as experimental controls.

INDIVIDUAL LABOR INCOME. — Table 4 reports the regression results for the IHS-transformed annual individual labor income of male respondents (Panel A) and female respondents (Panel B) on their own strategic thinking skills, following the three steps we outlined above. To save space, we do not report the coefficient estimates of the control variables here, but the full results are presented in Tables A6 and A7 in Appendix A.

We begin with male respondents. Columns (1) and (2) of Panel A report the baseline regression results for each measure of strategic thinking skills, controlling for sociodemographic variables. We find that a one-level increase in the BI score and a one-SD increase in the HOR score are associated with respective 42.9 percent and 68.1 percent increases in male respondents' annual labor income. The coefficient estimates are statistically significant at the 1 percent level.

The findings from the baseline regression might be partially driven by educational attainment, and cognitive ability, which can be correlated with both strategic thinking skills and labor income. Hence, in columns (3) and (4), we additionally control for educational attainment, IST score, and Eyes Test score. The coefficient estimates on the BI score and the HOR score—0.375 and 0.612, respectively—remain significant at the 5 percent level, although the magnitudes drop slightly by about 0.05 for the BI score and 0.07 for the HOR score.

Columns (5) and (6) show the regression results, further controlling for noncognitive traits and experimental variables. We find that a one-level increase in a male respondent's BI score is associated with a 37 percent increase in his annual income and a one-SD increase in his HOR score is associated with a 58 percent higher annual labor income. The magnitudes for the BI score and the HOR score decrease by only 13–15 percent compared with those from the baseline specification reported in columns (1) and (2).

We turn to female respondents. In the baseline specification, columns (1) and (2) of Panel B, we find that a one-level increase in a female respondent's BI score is associated with a 36.7 percent lower annual labor income. The coefficient

<sup>11</sup>We also investigate how much of the variation in labor market outcomes is explained by our measures of strategic thinking skills and the cognitive ability measures (IST score and Eyes Test Score). The findings are reported in Appendix A.A8.



Table 4—: Regression of individual labor income on strategic thinking skills

| Variables   | (1)                 | (2)                 | (3)                  | (4)                | (5)                  | (6)                |
|---|---------------------|---------------------|----------------------|--------------------|----------------------|--------------------|
| Dep. Var: IHS transformation of own annual labor income |                     |                     |                      |                    |                      |                    |
| Panel A: Male   |                     |                     |                      |                    |                      |                    |
| BI score  | 0.429***<br>(0.144) |                     | 0.375**<br>(0.149)   |                    | 0.372**<br>(0.151)   |                    |
| HOR score (standardized)                                |                     | 0.681***<br>(0.240) |                      | 0.612**<br>(0.248) |                      | 0.580**<br>(0.249) |
| Observations  | 1,044               | 1,044               | 1,044                | 1,044              | 1,044                | 1,044              |
| R-squared   | 0.037               | 0.037               | 0.041                | 0.041              | 0.053                | 0.052              |
| Panel B: Female   |                     |                     |                      |                    |                      |                    |
| BI score  | -0.367**<br>(0.179) |                     | -0.462***<br>(0.176) |                    | -0.473***<br>(0.182) |                    |
| HOR score (standardized)                                |                     | 0.257<br>(0.285)    |                      | -0.031<br>(0.289)  |                      | -0.006<br>(0.289)  |
| Observations  | 1,102               | 1,102               | 1,102                | 1,102              | 1,102                | 1,102              |
| R-squared   | 0.052               | 0.049               | 0.088                | 0.082              | 0.103                | 0.100              |
| Demographics  | Yes                 | Yes                 | Yes                  | Yes                | Yes                  | Yes                |
| Education and cognitive skills                          | No                  | No                  | Yes                  | Yes                | Yes                  | Yes                |
| Noncognitive and preference traits                      | No                  | No                  | No                   | No                 | Yes                  | Yes                |

*Note:* Standard errors are corrected for heteroskedasticity. \*\*\*, \*\*, \* denote  $p < 0.01$ ,  $p < 0.05$ ,  $p < 0.1$ , respectively. Columns (1)–(2) include only demographic variables: age group dummies, the ethnic Chinese dummy, marital status, number of children, spouse’s age, and the dummy variable reflecting a missing observation for spouse’s age for single individuals. Columns (3)–(4) additionally control for educational attainment, IST score, and Eyes Test score. Columns (5)–(6) additionally control for noncognitive traits such as financial planning time horizon, subjective risk tolerance, self-efficacy, personal optimism, and time taken to complete a corresponding task. Odd-numbered and even-numbered columns include dummy variables for the random orders of the Lift Game and the Line Game, respectively.

estimate is statistically significant at the 5 percent level. The coefficient estimate of the HOR score is positive but imprecisely estimated.

In columns (3) and (4), the negative association between the BI score and annual labor income for female respondents becomes even greater in its magnitude and significance after additionally controlling for educational attainment, the IST score, and the Eyes Test score. The association with the HOR level is close to zero and statistically insignificant.

In columns (5) and (6), with further controls for noncognitive and preference traits and experimental variables, the association of the BI score with female labor income remains statistically significant at the 1 percent level: a one-level increase in a female respondent’s BI score is associated with a 47 percent decrease in her annual labor income. The association with the HOR level becomes close to zero and statistically insignificant.

As we discussed before, the depth of reasoning exhibited by individuals in our

line game may not solely rely on their ability to engage in introspection but also on their beliefs about the behavior of their opponents. Thus, our HOR measure may capture some measurement error due to such beliefs. However, in the case of our BI measure, the experimental design mitigates this concern as subjects play against a computer player programmed to employ an optimal strategy in all scenarios. In this sense, the HOR measure could be more likely subject to measurement error compared to the BI measure, which attenuates the relation between female labor income and the HOR measure in Table 4.

In the spirit of [Altonji, Elder and Taber \(2005\)](#), the fact that the estimates are robust to a large extent to additional sets of controls supports the causal interpretation of the gender-dependent roles of strategic thinking skills in labor income. The negative association between the BI score and labor income for female respondents appears puzzling. However, Subsection [III.A](#) below shows that the gender-specific association between labor income and the BI score seems to be driven by the difference in extensive-margin labor supply decisions. In addition, we discuss potential channels to accommodate these findings in Section [IV](#).

Due to the nontrivial share of the participants reporting zero labor income, small changes at the low end of the income distribution may lead to disproportionate weights in the mean regression analysis. To address this concern, we report quantile regression analyses of the IHS-transformed labor income at different parts of the distribution in Tables [A9](#) and [A10](#) in Appendix [A](#). We find that the large magnitude reported in the mean regression analysis is indeed driven by disproportionately large effects at the low end of the income distribution, suggesting the role of strategic thinking skills in explaining the extensive margin of labor supply. Nevertheless, we also find that strategic thinking skills are related to labor income at the high end of the distribution as well. For male participants, a one-SD increase in the HOR score is associated with an increase of 15 percentage points and 12 percentage points at the 80th and 90th percentiles of the distribution, respectively. For female participants, a one-level increase of her BI score is associated with a decrease of 12 percentage points and 11 percentage points at the 80th and 90th percentiles of the distributions.

**EXTENSIVE MARGIN OF LABOR SUPPLY.** — The preceding analysis of labor income does not distinguish the extensive and intensive margins of labor supply because the IHS-transformed annual labor income contains observations with zero income. In this subsection, we examine the relationship between strategic thinking skills and the extensive margin of labor supply by gender.<sup>12</sup>

Columns (1)–(4) of Table [5](#) present the regression results for both male and

<sup>12</sup>The literature has documented that male and female labor supply has differential responsiveness to various factors, including the gender pay gap, marriage matching, cultural norms, and intrahousehold bargaining, compared with male labor supply ([Chiappori, 1992](#); [Pencavel, 1998](#); [Blundell et al., 2007](#); [Chiappori and Mazzocco, 2017](#))

female respondents’ labor supply along the extensive margin on their strategic thinking skills, including the full set of controls. The dependent variable takes the value of 1 if a respondent has a positive annual labor income and 0 otherwise. This measure includes both working for pay and self-employment.

Columns (1)–(2) report that male respondents with higher values of the BI and HOR scores are more likely to earn positive labor income. A one-level increase in a male respondent’s BI score is associated with an increase of 3 percentage points in the probability of his employment, and a one-SD increase in his HOR score is associated with an increase of 4.7 percentage points. Both estimates are statistically significant at the 5 percent level.

Columns (3)–(4) show that, in contrast to the case of male respondents, a female respondent’s BI score is negatively correlated with the likelihood of earning a positive annual labor income. A one-level increase in a female respondent’s BI score is associated with a decrease of 4.2 percentage points in the probability of a positive annual labor income at the 5 percent significance level.

Table 5—: Regression results for the extensive margin of labor supply by gender

| Variables                | (1)                         | (2)                | (3)                 | (4)              | (5)                   | (6)               | (7)                 | (8)              |
|--------------------------|-----------------------------|--------------------|---------------------|------------------|-----------------------|-------------------|---------------------|------------------|
|                          | I (Annual labor income > 0) |                    |                     |                  | Retired or Unemployed |                   | Homemaker           |                  |
|                          | Male                        |                    | Female              |                  | Male                  |                   | Female              |                  |
| BI score                 | 0.030**<br>(0.013)          |                    | -0.042**<br>(0.017) |                  | -0.035**<br>(0.014)   |                   | 0.053***<br>(0.018) |                  |
| HOR score (standardized) |                             | 0.047**<br>(0.022) |                     | 0.002<br>(0.027) |                       | -0.035<br>(0.022) |                     | 0.037<br>(0.031) |
| Observations             | 1,044                       | 1,044              | 1,102               | 1,102            | 1,044                 | 1,044             | 821                 | 821              |
| R-squared                | 0.044                       | 0.042              | 0.082               | 0.079            | 0.103                 | 0.098             | 0.064               | 0.059            |

*Note:* Standard errors corrected for heteroskedasticity are reported in parentheses. All columns include age group dummies, the ethnic Chinese dummy, marital status, number of children, spouse’s age, and the dummy variable reflecting a missing observation for spouse’s age for single individuals, educational attainment, IST score, Eyes Test score, financial planning time horizon, subjective risk tolerance, self-efficacy, personal optimism, and the time taken to complete each task. Odd-numbered and even-numbered columns include dummy variables for the random orders of the Lift Game and the Line Game, respectively. \*\*\*, \*\*, \* denote  $p < 0.01$ ,  $p < 0.05$ ,  $p < 0.1$ , respectively.

These results suggest that the gender-dependent associations between labor income and strategic thinking skills documented in Subsection III.A are driven at least partly by gender differences in the association between the extensive margin decision of labor supply and strategic thinking skills.<sup>13</sup>

Where does the gender difference in the associations between strategic thinking skills and the extensive margin of labor supply come from? To answer this question, we examine the composition of self-reported labor status among the

<sup>13</sup>In Appendix A.A9, we conduct an analysis of intensive margin labor supply decisions by shutting down the extensive margin channel (i.e., by excluding zero-income earners). The regression results reported in Table A11 indicate that the associations between strategic thinking skills and individual labor incomes are in the same direction but are imprecisely estimated when we exclude those who report zero labor income.

respondents who reported zero annual labor income. Figure 1 shows a stark gender difference in terms of the labor status composition for zero-income earners. Among male respondents, retirees and the unemployed account for the majority of zero-income earners, while among female respondents, the majority of zero-income earners report themselves as homemakers.<sup>14</sup>

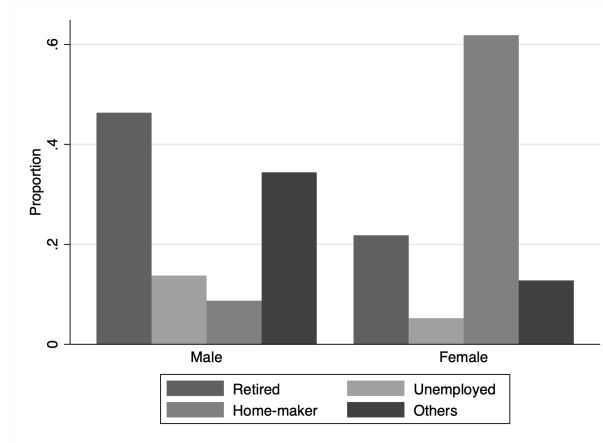


Figure 1. : Labor status composition of participants with zero labor income

To further examine whether the gender differences in the composition of labor status contribute to the gender differences in the association between strategic thinking skills and the extensive margin of labor supply, we conduct regression analyses of either being retired or unemployed for males and the homemaker status for females on the strategic thinking skill measures while including the full set of controls.

Columns (5)–(6) of Table 5 report the regression results for male respondents' labor status of being either retired or unemployed on strategic thinking skills with the full set of controls. The dependent variable takes the value of 1 if a male respondent declares himself to be retired or unemployed and 0 otherwise. The results show that male respondents with higher values of the BI and HOR scores are less likely to be retired or unemployed. Overall, the negative associations between strategic thinking skills and zero income due to retirement and unemployment are consistent with the earlier findings, reported in Panel A of Table 4 and Columns (1)–(2) of Table 5.

Columns (7)–(8) of Table 5 report the regression results for female respondents. The dependent variable takes the value of 1 if a female respondent declares herself to be a homemaker and 0 otherwise. We restrict the sample of female respondents in this analysis to those who were married at the time of the survey. In column

<sup>14</sup>The "others" category includes studying, disability, sick leave, etc.

(7), a one-level increase in the married female respondent's BI score is associated with a 5.3 percent increase in the likelihood of being a homemaker. This evidence is consistent with the earlier findings that a female respondent's BI score is negatively correlated with the likelihood of working and, as a result, with her annual labor income. In column (8), we find a positive association between a married female respondent's HOR score and her homemaker status, but the estimate is not statistically significant.

We believe that the finding that only the BI measure but not the HOR measure exhibits statistical significance in the extensive margin of labor supply is sensible. The nature of this extensive margin decision requires individuals to carefully compare outcomes across multiple future contingencies, such as both household members working and earning labor income, only the husband working, or only the wife working. Particularly for those females who decide to specialize in home production, it is necessary to compare the future flow of costs and benefits associated with these different contingencies.

### *B. Strategic Thinking and Household Labor Income*

The gender-dependent relationship with individual labor outcomes raises a question about whether strategic thinking skills play a role beyond the scope of individual outcomes. We address this question by examining the association between household labor income and married respondents' strategic thinking skills. Table 6 reports the regression results for the IHS-transformed household annual labor income for male married respondents (columns (1) and (2)), for female married respondents (columns (3) and (4)) and for female married respondents whose individual labor income is zero (columns (5) and (6)). All reported results include the full set of controls.

Regardless of gender, we find significantly positive associations between household labor income and individual-level strategic thinking skills. A one-SD increase in the HOR score is associated with a 60.6 percent increase in household labor income for male married respondents and a 62.1 percent increase for female married respondents. The estimates are significant at the one percent level for males and at the five percent level for females. Similarly, a one-level increase in the BI score is associated with a 6.4 percent increase in household labor income for male respondents and a 14.4 percent increase for female married respondents, although the estimates are imprecisely estimated.

Since females are less likely to be working in our data, we further examine the association for married females whose individual labor income is zero. The results reported in columns (5) and (6) indicate that a one-level increase in the BI score and a one-SD increase in the HOR score are associated with an 87.3 percent increase at the one percent significance level and a 113.6 percent increase in household labor income at the five percent significance level, respectively.

We find that, unlike the gender-dependent associations with individual labor income, strategic thinking skills are positively associated with household labor

income independently of gender. This finding suggests the potential importance of strategic thinking skills for collective economic success at the household level.

Table 6—: Regression of household labor income on married respondents’ strategic thinking skills

| Variables                | (1)<br>Male      | (2)                 | (3)<br>Female    | (4)                | (5)<br>Non-working Female | (6)                |
|--------------------------|------------------|---------------------|------------------|--------------------|---------------------------|--------------------|
| BI score                 | 0.064<br>(0.125) |                     | 0.144<br>(0.135) |                    | 0.873***<br>(0.315)       |                    |
| HOR score (standardized) |                  | 0.606***<br>(0.212) |                  | 0.621**<br>(0.261) |                           | 1.136**<br>(0.506) |
| Observations             | 938              | 938                 | 822              | 822                | 338                       | 338                |
| R-squared                | 0.047            | 0.055               | 0.148            | 0.150              | 0.209                     | 0.200              |

*Note:* Standard errors corrected for heteroskedasticity are reported in parentheses. All columns include age group dummies, the ethnic Chinese dummy, marital status, number of children, spouse’s age, the dummy variable reflecting a missing observation for spouse’s age for single individuals, educational attainment, IST score, Eyes Test score, financial planning time horizon, subjective risk tolerance, self-efficacy, personal optimism, and the time taken to complete each task. Odd-numbered and even-numbered columns include dummy variables for the random orders of the Lift Game and the Line Game, respectively. \*\*\*, \*\*, and \* denote  $p < 0.01$ ,  $p < 0.05$ , and  $p < 0.1$ , respectively.

#### IV. Discussion on Potential Channels

What drives the observed household-level, gender-independent, and individual-level, gender-dependent associations between strategic thinking skills and labor market outcomes? We consider broadly two potential channels: *labor market channels* and *household interaction channels*.

##### A. Labor Market Channels

We first explore whether the labor market channel through which individuals’ strategic thinking skills affect their labor and household incomes can provide a coherent account of the empirical findings in this paper. Two potential channels are worth considering: a workplace channel and an occupation-choice channel. The former suggests that individuals with higher strategic thinking skills may contribute more effectively to team production in the workplace and hence earn higher wages. This idea has been studied by Deming (2017), who proposes that social skills distinct from cognitive skills can enhance labor productivity by facilitating task specialization. While this channel can explain the positive association between strategic thinking skills and labor income among male respondents, it does not account for the opposite pattern found among female respondents, nor does it explain gender-dependent patterns regarding the extensive margin of labor supply. Therefore, this channel may have limited explanatory power for our main empirical findings.

We turn to the occupation-choice channel whereby individuals with higher strategic thinking skills are sorted into professions and industries with higher demand for these skills. The SLP collected information about the respondent's occupation as a special module of the July 2017 wave and thus we examine whether their past career choices are associated with strategic thinking skills. It asked the respondent's longest-held, last (if not working), and current (if working) occupation using the Singapore Standard Occupation Classification code. To analyze the association between strategic thinking skills and respondents' past occupation choice, we constructed the following two dependent variables: (i) whether a respondent's longest-held job was a senior manager, professional, and sales and service worker and (ii) the social skill score of the longest-held job. We assume that the occupation categories of the first variable are likely to require a higher level of strategic thinking skills. The latter variable is computed using the U.S. Department of Labor's O\*NET database, which provides the skill requirements associated with each occupation following [Deming \(2017\)](#).<sup>15</sup>

Table A19 in Appendix A reports the regression results of each dependent variable by gender. We find that our strategic thinking skill measures are not significantly associated with the level of social skills required by respondents' longest-held jobs. The results remain robust when using the information on current or last occupation. With the caution that the sample size in this analysis is smaller than the baseline sample because the occupation information was collected as part of a special module, we interpret that these results provide little evidence supporting the occupation-choice channel for our data.

In summary, we conclude that the labor market channels alone cannot provide a coherent account of both the household-level gender-independent associations and individual-level gender-dependent associations between strategic thinking skills and labor market outcomes.

### B. *Intrahousehold Interaction Channels*

Next, we investigate whether the intrahousehold interaction channel through which individuals' strategic thinking skills play a role in intrahousehold decision making which further affects their labor and household incomes can provide a coherent account of our main empirical findings. There are two plausible such channels: i) an intrahousehold decision-making of labor supply with home-to-workplace spillover, and ii) an assortative matching on strategic thinking skills in marriage. The former suggests that when the division of labor for domestic and marketable good production occurs between a wife and her husband based on their comparative advantage or gender norms, strategic thinking skills can facilitate better coordination and improve the spillover from home to the workplace. The positive spillover from home to the workplace is well documented in

<sup>15</sup>We first cross-walked the U.S. Standard Occupation Classification code and that of Singapore and merged O\*NET's occupation. A higher value indicates that the occupation requires more social skills. See [Deming \(2017\)](#) for details.



the psychology literature (Barnett, 1994; Barnett and Marshall, 1992*a,b*; Kirchmeyer, 1992) as well as in the economics literature (Benham, 1974; Lefgren and McIntyre, 2006; Huang et al., 2009). It is plausible that factors such as higher quality meals, better care for children, social status, leisure activities, companionship, love, and improved health can enhance an individual's productivity in the labor market.

In Appendix A.A11, we propose a model of household labor supply built upon the literature on collective labor supply with household production and workplace production (e.g., Apps and Rees, 1997; Chiappori, 1997). Following the literature, a household consisting of two individuals achieves a Pareto-efficient resource allocation between market labor supply and time spent in domestic production under the condition that the domestic good increases the productivity of the marketable good production. Individuals who are heterogeneous with respect to their productivity in producing domestic goods and marketable goods need to coordinate intrahousehold specialization between these two production tasks. Strategic thinking skills are assumed to facilitate coordination between them and create a larger degree of home-to-workplace spillover. It is a natural adoption of the production function introduced in the literature on task allocation in the workplace (e.g., Autor, Levy and Murnane, 2003; Acemoglu and Autor, 2011; Autor and Handel, 2013; Deming, 2017).

This model of collective labor supply offers a set of predictions consistent with the following empirical findings in our data, while it is silent about the high correlation of strategic thinking skills within the married couple. First, both married males' and females' strategic thinking skills are positively associated with their household labor income. Second, males' strategic thinking skills are positively associated with their individual labor income and negatively associated with the likelihood of being retired; females' strategic thinking skills are negatively associated with their individual labor income and positively associated with the likelihood of being homemakers. Third, as reported in Table 6, we also found that married, non-market-participating females' strategic thinking skills are positively associated with household labor income. Fourth, as reported in Appendix A.A14, females' strategic thinking skills are positively associated with their spouses' individual labor income.

An alternative possibility in the household interaction channels is assortative matching in marriage, whereby individuals marry someone with similar strategic thinking skills. This channel assumes that strategic thinking is socially valuable, and therefore, individuals seek partners who possess these skills. This channel offers a couple of insights. First, strategic thinking skills between the couple members are positively correlated as we reported in Section II. Second, if individual strategic thinking skills are positively associated with the traditional sources of determining labor outcomes such as education and cognitive ability (as we found in Table 3), this channel may generate the positive relation between strategic thinking skills and individual and household labor incomes. However, this

channel alone is unable to explain why some couples specialize between household production and workplace production nor why individual income resulting from this specialization is correlated with strategic thinking skills in a gender-dependent manner.<sup>16</sup>

Therefore, we conclude that the intrahousehold interaction channels combining the collective model of labor supply and the assortative marriage matching channel can accommodate the key empirical findings of this paper.

## V. Robustness Check with ORIV Approach

One potential concern regarding the analysis of the paper is potential measurement error in eliciting strategic thinking skills and the lack of corrections of any noise in these elicitations in establishing the relation between strategic thinking skills and labor outcomes. One form of this concern can be expressed as that our HOR and BI measures are just two experimental proxies of a common trait of strategic thinking and are measured with errors. We take this concern for granted and attempt to remove the attenuation bias caused by potential measurement error by following [Gillen, Snowberg and Yariv \(2019\)](#)'s obviously related instrumental variables (ORIV) method. It is worth emphasizing that the ORIV approach assumes that BI and HOR measures capture the same underlying trait (strategic thinking skill) with idiosyncratic errors. In contrast, as we addressed in Section II, we argue that the nature of strategic thinking is multi-dimensional and our HOR and BI measures capture two distinct features of strategic thinking.

Table 7 reports the ORIV estimation results on the relation between strategic thinking skills and individual and household labor income for male and female participants, respectively. To make the scale consistent, we standardized the BI score using a normal distribution as in the HOR score. Then we assume that the measurement errors in BI and HOR scores are uncorrelated with one another. Banking on this assumption, we use BI as an instrument for HOR and vice versa to remove measurement error and estimated the two-stage least squares regression where strategic thinking skill is the endogenous variable with measurement error in the second stage.<sup>17</sup>

The ORIV estimation results are consistent with the gender-dependent associations with individual labor income and the gender-independent associations with

<sup>16</sup>The marriage assortative matching channel suggests that a non-working household member's higher strategic thinking skill contributes to increased household income solely because they selected a partner with higher strategic thinking skills. To investigate this, we conducted a regression analysis, additionally controlling for the strategic thinking skill of the working spouse, using a sample of 138 married couples where only one household member is active in the labor market (see Table A20 in Appendix A). Despite the caveat regarding the limited sample size and the generalizability of these findings, we observed that the non-working spouse's strategic thinking skill in terms of the HOR measure remained significant but the estimates are not statistically significant.

<sup>17</sup>Following [Gillen, Snowberg and Yariv \(2019\)](#), we stacked the duplicated data to further improve statistical precision. The assumption of this approach is that the underlying strategic thinking skill (measured with error) is BI (HOR) score and the IV is HOR (BI) score in the original (duplicated) dataset.

household labor income in the previous analysis assuming that our measures of HOR and BI capture two distinct features of strategic thinking skills. For male participants, the strategic thinking skill is positively associated with their individual income at the 1% significant level. However, it is negatively correlated with females' individual income, and the estimate is statistically significant at the 5% level. In addition, we find consistent evidence that the strategic thinking skill can increase household labor income regardless of gender, as shown in Table 6. The estimates are statistically significant at 5% or 10% level. Hence, we conclude that the key findings with individual and household labor incomes are immune from the concerns of measurement errors.

Table 7—: ORIV estimation of labor income on strategic thinking skills

|                          | (1)                     | (2)                 | (3)                           | (4)                |
|--------------------------|-------------------------|---------------------|-------------------------------|--------------------|
|                          | IHS of own labor income |                     | IHS of household labor income |                    |
|                          | Male                    | Female              | Male                          | Female             |
| Strategic thinking skill | 1.703***<br>(0.608)     | -1.782**<br>(0.890) | 0.781*<br>(0.472)             | 1.470**<br>(0.707) |
| Observations             | 2,088                   | 2,204               | 1,876                         | 1,644              |
| R-squared                | 0.009                   | 0.054               | 0.041                         | 0.102              |

*Note:* Standard errors clustered at the respondent level are reported in parentheses. All columns include age group dummies, the ethnic Chinese dummy, marital status, number of children, spouse's age, the dummy variable reflecting a missing observation for spouse's age for single individuals, educational attainment, IST score, the Eyes Test score, financial planning time horizon, subjective risk tolerance, self-efficacy, personal optimism, random orders of the Lift Game and Line Game, and time taken to complete each task. \*\*\*, \*\*, \* denote  $p < 0.01$ ,  $p < 0.05$ ,  $p < 0.1$ , respectively.

## VI. Concluding Remarks

In our large-scale experiment, we measured strategic thinking skills, specifically focusing on fundamental aspects of strategic reasoning in interpersonal interaction. These measures, capturing higher-order rationality and backward induction, were found to be distinct from the conventional collection of cognitive and noncognitive skills. Notably, both married males and married females exhibited significant and positive associations between their strategic thinking skills and their household labor income, emphasizing the importance of strategic thinking skills for collective economic success within a household. However, we also observed contrasting roles of strategic thinking skills across genders in their individual labor market outcomes. Males' strategic thinking skills were positively associated with both individual labor income and labor market participation, while females' strategic thinking skills showed a negative association with these outcomes. We propose intrahousehold interaction channels as a potential explanation for the

patterns observed in our study. In sum, our findings strongly support the notion that strategic thinking is a skill of economic significance.

One might be concerned that our findings are obtained from older adults in Singapore and thus its external validity is limited due to the differences in cultural norms and age. In Appendix A.A17, we provide suggestive evidence that our main findings are unlikely to change significantly.

As acknowledged earlier, the two strategic thinking measures developed in our paper only cover elementary aspects of strategic thinking skills, and there are likely several other aspects that are relevant to real-life decision-making. One such aspect is related to subjects' learning rules and the speed at which they learn, as our game-theoretic measures of strategic thinking skills focus on subjects' initial play, without considering learning opportunities. The experimental literature on learning in games has demonstrated heterogeneity among human subjects in their sophistication when it comes to learning rules (e.g., Camerer and Ho, 1999). The capacity to effectively learn in strategic environments may be associated with individuals' economic and social success. Future research could explore the development of a measure for learning capabilities and investigate its relationship with economic performance. Additionally, there is a need for further exploration of scientific interventions aimed at enhancing strategic thinking skills and ultimately improving economic outcomes.

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